

Weierstrass Institute for Applied Analysis and Stochastics



Computational finance – Lecture 3

Christian Bayer

Mohrenstrasse 39 · 10117 Berlin · Germany · Tel. +49 30 20372 0 · www.wias-berlin.de



1 Variance reduction

2 Quasi Monte Carlo

Computational finance – Lecture 3 · April 30, 2021 · Page 2 (14)



Antithetic variates



• $U \sim \mathcal{U}(]0, 1[) \Rightarrow \widetilde{U} \coloneqq 1 - U \sim \mathcal{U}(]0, 1[), W \sim \mathcal{N}(0, I_d) \Rightarrow \widetilde{W} \coloneqq -W \sim \mathcal{N}(0, I_d).$ In an intuitive sense, in both cases *X* is anti-monotonic to \widetilde{X} .

Definition (Antithetic variates)

Given a r.v. X and a (deterministic) transformation \widetilde{X} with the same distribution.

$$I_M^{\mathbf{A}}[f; X, \widetilde{X}] \coloneqq \frac{1}{M} \sum_{i=1}^M \frac{f(X_i) + f(\widetilde{X}_i)}{2}.$$

Antithetic variates



► $U \sim \mathcal{U}(]0,1[) \Rightarrow \widetilde{U} \coloneqq 1 - U \sim \mathcal{U}(]0,1[), W \sim \mathcal{N}(0,I_d) \Rightarrow \widetilde{W} \coloneqq -W \sim \mathcal{N}(0,I_d).$ In an intuitive sense, in both cases *X* is anti-monotonic to \widetilde{X} .

Definition (Antithetic variates)

Given a r.v. X and a (deterministic) transformation \widetilde{X} with the same distribution.

$$I_M^{\mathbf{A}}[f; X, \widetilde{X}] \coloneqq \frac{1}{M} \sum_{i=1}^M \frac{f(X_i) + f(\widetilde{X}_i)}{2}.$$

- Assume that cost(X, X) ≤ 2 cost(X), i.e., cost of simulation of X and transformation to X is at most twice the cost of simulation of X.
- ▶ Then antithetic variates are more efficient than standard MC iff $\operatorname{cov}(f(X), f(\widetilde{X})) < 0$.



Control variates



Suppose we are given a r.v. Y and a function g s.t. I[g; Y] is known. Then, for $\lambda \in \mathbb{R}$,

$$I[f;X] = E[f(X) - \lambda g(Y)] + \lambda I[g;Y].$$

Definition (Control variates)

$$I_M^{\boldsymbol{C},\lambda}[f;X,Y] \coloneqq \frac{1}{M} \sum_{i=1}^M \left(f(X_i) - \lambda g(Y_i)\right) + \lambda I[g;Y].$$





Suppose we are given a r.v. Y and a function g s.t. I[g; Y] is known. Then, for $\lambda \in \mathbb{R}$,

$$I[f;X] = E\left[f(X) - \lambda g(Y)\right] + \lambda I[g;Y].$$

Definition (Control variates)

$$I_M^{\boldsymbol{C},\lambda}[f;X,Y] \coloneqq \frac{1}{M} \sum_{i=1}^M \left(f(X_i) - \lambda g(Y_i) \right) + \lambda I[g;Y].$$

- Often Y = X, and the cost of $I_M^{C,\lambda}$ is only insignificantly higher than the cost of I_M .
- The MSE of $I_M^{C,\lambda}$ is minimized by $\lambda^* := \frac{\operatorname{cov}(f(X),g(Y))}{\operatorname{var}(g(Y))}$, the minimizer of $\operatorname{var}(f(X) \lambda g(Y))$.
- $MSE[I_M^{C,\lambda^*}] = (1 \rho^2) MSE[I_M]$, where $\rho := cor(f(X), g(Y)) \Rightarrow |\rho|$ as large as possible!





Suppose we are given a r.v. Y and a function g s.t. I[g; Y] is known. Then, for $\lambda \in \mathbb{R}$,

$$I[f;X] = E\left[f(X) - \lambda g(Y)\right] + \lambda I[g;Y].$$

Definition (Control variates)

$$I_M^{\boldsymbol{C},\lambda}[f;X,Y] \coloneqq \frac{1}{M} \sum_{i=1}^M \left(f(X_i) - \lambda g(Y_i) \right) + \lambda I[g;Y].$$

- Often Y = X, and the cost of $I_M^{C,\lambda}$ is only insignificantly higher than the cost of I_M .
- The MSE of $I_M^{C,\lambda}$ is minimized by $\lambda^* := \frac{\operatorname{cov}(f(X),g(Y))}{\operatorname{var}(g(Y))}$, the minimizer of $\operatorname{var}(f(X) \lambda g(Y))$.
- MSE $[I_M^{C,\lambda^*}] = (1 \rho^2)$ MSE $[I_M]$, where $\rho := \operatorname{cor}(f(X), g(Y)) \Rightarrow |\rho|$ as large as possible!

Example

 $E[S_T] = S_0$, hence $Y = S_T$, g = id can be used as control variate for option pricing.





- Black Scholes model (r = 0 as usual): $S_t = S_0 \exp(\sigma W_t \frac{1}{2}\sigma^2 t)$.
- Asian option: option depending on the time average of the stock price, e.g.,

$$f(X) \coloneqq \left(\frac{1}{N}\sum_{n=1}^{N} X_n - K\right)^+, \quad X \coloneqq (S_{t_1}, \ldots, S_{t_N}), \quad 0 \le t_1 < \cdots < t_N \le T.$$



- Black Scholes model (r = 0 as usual): $S_t = S_0 \exp(\sigma W_t \frac{1}{2}\sigma^2 t)$.
- Asian option: option depending on the time average of the stock price, e.g.,

$$f(X) := \left(\frac{1}{N} \sum_{n=1}^{N} X_n - K\right)^+, \quad X := (S_{t_1}, \dots, S_{t_N}), \quad 0 \le t_1 < \dots < t_N \le T.$$

No explicit price formula as sum of log-normal r.v.s are not log-normal.



- Black Scholes model (r = 0 as usual): $S_t = S_0 \exp(\sigma W_t \frac{1}{2}\sigma^2 t)$.
- Asian option: option depending on the time average of the stock price, e.g.,

$$f(X) \coloneqq \left(\frac{1}{N}\sum_{n=1}^{N} X_n - K\right)^+, \quad X \coloneqq (S_{t_1}, \dots, S_{t_N}), \quad 0 \le t_1 < \dots < t_N \le T.$$

- No explicit price formula as sum of log-normal r.v.s are not log-normal.
- MC simulation: For $Z \sim \mathcal{N}(0, I_N)$, set $W_{t_1} = \sqrt{t_1}Z_1$ and $W_{t_n} = W_{t_{n-1}} + \sqrt{t_n t_{n-1}}Z_n$.



- Black Scholes model (r = 0 as usual): $S_t = S_0 \exp(\sigma W_t \frac{1}{2}\sigma^2 t)$.
- Asian option: option depending on the time average of the stock price, e.g.,

$$f(X) \coloneqq \left(\frac{1}{N}\sum_{n=1}^{N} X_n - K\right)^+, \quad X \coloneqq (S_{t_1}, \dots, S_{t_N}), \quad 0 \le t_1 < \dots < t_N \le T.$$

- No explicit price formula as sum of log-normal r.v.s are not log-normal.
- MC simulation: For $Z \sim \mathcal{N}(0, I_N)$, set $W_{t_1} = \sqrt{t_1}Z_1$ and $W_{t_n} = W_{t_{n-1}} + \sqrt{t_n t_{n-1}}Z_n$.
- Observation: $\frac{1}{N} \sum_{n=1}^{N} x_n \approx \left(\prod_{n=1}^{N} x_n \right)^{1/N}$ (arithmetic vs. geometric mean), and

$$\left(\prod_{n=1}^{N} S_{t_n}\right)^{1/N} = S_0 \exp\left(\frac{\sigma}{N} \sum_{n=1}^{N} W_{t_n} - \frac{\sigma^2}{2N} \sum_{n=1}^{N} t_n\right) \sim \operatorname{LN}\left(\log S_0 - \frac{\sigma^2}{2N} \sum_{n=1}^{N} t_n, \ \frac{\sigma^2}{N^2} \sum_{n=1}^{N} (1 + 2(N - n)) t_n\right).$$





- Black Scholes model (r = 0 as usual): $S_t = S_0 \exp(\sigma W_t \frac{1}{2}\sigma^2 t)$.
- Asian option: option depending on the time average of the stock price, e.g.,

$$f(X) \coloneqq \left(\frac{1}{N}\sum_{n=1}^{N} X_n - K\right)^+, \quad X \coloneqq \left(S_{t_1}, \ldots, S_{t_N}\right), \quad 0 \le t_1 < \cdots < t_N \le T.$$

- No explicit price formula as sum of log-normal r.v.s are not log-normal.
- MC simulation: For $Z \sim \mathcal{N}(0, I_N)$, set $W_{t_1} = \sqrt{t_1}Z_1$ and $W_{t_n} = W_{t_{n-1}} + \sqrt{t_n t_{n-1}}Z_n$.
- Observation: $\frac{1}{N} \sum_{n=1}^{N} x_n \approx \left(\prod_{n=1}^{N} x_n\right)^{1/N}$ (arithmetic vs. geometric mean), and

$$\left(\prod_{n=1}^{N} S_{t_n}\right)^{1/N} = S_0 \exp\left(\frac{\sigma}{N} \sum_{n=1}^{N} W_{t_n} - \frac{\sigma^2}{2N} \sum_{n=1}^{N} t_n\right) \sim \operatorname{LN}\left(\log S_0 - \frac{\sigma^2}{2N} \sum_{n=1}^{N} t_n, \frac{\sigma^2}{N^2} \sum_{n=1}^{N} (1 + 2(N - n)) t_n\right).$$

By the Black-Scholes formula, a control variate is given by the geometric Asian option

$$g(Y) := \left(\left(\prod_{n=1}^{N} Y_n \right)^{1/N} - K \right)^+, \quad Y := X = \left(S_{t_1}, \ldots, S_{t_N} \right).$$





- Idea: Sample more where the "local variance" is higher.
- Setting: X has (d-dimensional) density p, additional r.v. Y with density q.





Setting: *X* has (*d*-dimensional) density *p*, additional r.v. *Y* with density *q*.

Definition (Importance sampling)

$$I_M^{IS}[f;X,Y] \coloneqq \frac{1}{M} \sum_{i=1}^M f(Y_i) \frac{p(Y_i)}{q(Y_i)} = I_M\left[f\frac{p}{q};Y\right].$$





Setting: *X* has (*d*-dimensional) density *p*, additional r.v. *Y* with density *q*.

Definition (Importance sampling)

$$I_M^{IS}[f;X,Y] \coloneqq \frac{1}{M} \sum_{i=1}^M f(Y_i) \frac{p(Y_i)}{q(Y_i)} = I_M\left[f\frac{p}{q};Y\right].$$

Mean squared error determined by

$$\operatorname{var}\left(f(Y)\frac{p(Y)}{q(Y)}\right) = E\left[f(Y)^2\frac{p(Y)^2}{q(Y)^2}\right] - I[f;X]^2 = E\left[f(X)^2\frac{p(X)}{q(X)}\right] - I[f;X]^2.$$



- Idea: Sample more where the "local variance" is higher.
- Setting: X has (d-dimensional) density p, additional r.v. Y with density q.

Definition (Importance sampling)

$$I_M^{IS}[f;X,Y] \coloneqq \frac{1}{M} \sum_{i=1}^M f(Y_i) \frac{p(Y_i)}{q(Y_i)} = I_M\left[f\frac{p}{q};Y\right].$$

Mean squared error determined by

$$\operatorname{var}\left(f(Y)\frac{p(Y)}{q(Y)}\right) = E\left[f(Y)^2\frac{p(Y)^2}{q(Y)^2}\right] - I[f;X]^2 = E\left[f(X)^2\frac{p(X)}{q(X)}\right] - I[f;X]^2.$$

▶ Best possible speed-up: *q* proportional to $f \cdot p$, implying $f(Y)\frac{p(Y)}{q(Y)} \equiv 1$ – but requires knowledge of I[f; X] for normalization.





1 Variance reduction

2 Quasi Monte Carlo

Computational finance - Lecture 3 · April 30, 2021 · Page 7 (14)



More evenly distributed numbers





Computational finance – Lecture 3 · April 30, 2021 · Page 8 (14)



Discrepancy



- Intuition: discrepancy is the quadrature error for indocator functions of rectangles R.
- Consider point sets $(x_i)_{i=1}^M \subset (x_i)_{i\in\mathbb{N}}$ or $(x_i^M)_{i=1}^M$, $x_i \in [0, 1]^d$.

Discrepancy

D



- Intuition: discrepancy is the quadrature error for indocator functions of rectangles R.
- Consider point sets $(x_i)_{i=1}^M \subset (x_i)_{i\in\mathbb{N}}$ or $(x_i^M)_{i=1}^M$, $x_i \in [0, 1]^d$.

Definition (Discrepancy)

The discrepancy D_M and star-discrepancy D_M^* for a point set $(x_i)_{i=1}^M$ are defined as

$$D_{M} = \sup_{\substack{\text{rectangles } R \subset [0,1]^{d} \\ M}} \left| \frac{1}{M} \# \{ 1 \le i \le M : x_{i} \in R \} - \lambda(R) \right|,$$

$$M_{M} = \sup \left\{ \left| \frac{1}{M} \# \{ 1 \le i \le M : x_{i} \in R \} - \lambda(R) \right| : R = \sum_{j=1}^{d} [0, b_{j}[, b_{1}, \dots, b_{d} \in [0, 1] \}.$$



Discrepancy

 D_M^*



- Intuition: discrepancy is the quadrature error for indocator functions of rectangles R.
- Consider point sets $(x_i)_{i=1}^M \subset (x_i)_{i\in\mathbb{N}}$ or $(x_i^M)_{i=1}^M, x_i \in [0, 1]^d$.

Definition (Discrepancy)

The discrepancy D_M and star-discrepancy D_M^* for a point set $(x_i)_{i=1}^M$ are defined as

$$D_{M} = \sup_{\text{rectangles } R \subset [0,1]^{d}} \left| \frac{1}{M} \# \{ 1 \le i \le M : x_{i} \in R \} - \lambda(R) \right|,$$

= $\sup \left\{ \left| \frac{1}{M} \# \{ 1 \le i \le M : x_{i} \in R \} - \lambda(R) \right| : R = \sum_{j=1}^{d} [0, b_{j}[, b_{1}, \dots, b_{d} \in [0, 1] \} \right\}$

Definition (Low discrepancy)

A sequence of point sets $(x_i^M)_{i=1}^M$, $M \in \mathbb{N}$, has low discrepancy iff $D_M^* \leq c \frac{\log(M)^d}{M}$.



Variation



► Regularity of integrand *f* classically measured by variation.

Definition (Variation in the sense of Hardy-Krause)

For a one-dimensional function $f:[0,1] \to \mathbb{R}$

$$V[f] \coloneqq \int_0^1 \left| \frac{df}{dx}(x) \right| dx$$

and for $f: [0,1]^d \to \mathbb{R}$

$$V[f] \coloneqq \int_{[0,1]^d} \left| \frac{\partial^d f}{\partial x^1 \cdots \partial x^d}(x) \right| dx + \sum_{j=1}^d V[f_1^{(j)}],$$

where $f_1^{(j)}$ denotes the restriction of *f* to the boundary $x^j = 1$.

Computational finance – Lecture 3 · April 30, 2021 · Page 10 (14)





Theorem (Koksma-Hlawka inequality)

 $|I[f] - J_M[f]| \le V[f]D_M^*.$







Theorem (Koksma-Hlawka inequality)

 $|I[f] - J_M[f]| \le V[f]D_M^*.$

- Error factorization into term depending on regularity of integrand and uniformity of point set.
- In contrast to MC error, the Koksma-Hlawka inequality is deterministic, but usually far from sharp! (Note that it essentially require *f* to be C^d!)



Theorem (Koksma-Hlawka inequality)

 $|I[f] - J_M[f]| \le V[f]D_M^*.$

- Error factorization into term depending on regularity of integrand and uniformity of point set.
- In contrast to MC error, the Koksma-Hlawka inequality is deterministic, but usually far from sharp! (Note that it essentially require *f* to be C^d!)

Example (Van der Corput sequence)

Let *p* prime and for $k \in \mathbb{N}_0$ let $(a_j(k))_{j \in \mathbb{N}_0}$ denote the p-ary expansion of *k*, i.e., $k = \sum_{i=0}^{\infty} a_i(k) p^i$. A one-dimensional low-discrepancy sequence is given by

$$x_i \coloneqq \psi_p(i) \coloneqq \sum_{j=0}^{\infty} \frac{a_j(i)}{p^{j+1}}, \quad i \in \mathbb{N}_0.$$



Numerical comparison





Figure: A call option in the Black-Scholes model using Monte Carlo and Quasi Monte Carlo simulation, Red: MC simulation, blue: QMC simulation, black: **Reference lines** proportional to 1/Mand $1/\sqrt{M}$.



Computational finance - Lecture 3 · April 30, 2021 · Page 12 (14)

- For uniform (pseudo) r.v.s: If $U_i \sim \mathcal{U}([0, 1])$, (U_i) independent, then $(U_{(i-1)d+1}, \ldots, U_{id})_{i \in \mathbb{N}}$ is a sequence of independent r.v.s from $\mathcal{U}([0, 1]^d)$.
- In contrast, if (x_i)_{i∈ℕ} is one-dimensional low-discrepancy sequence, then (x_{(i-1)d+1},..., x_{id})_{i∈ℕ} is most likely not a d-dimensional low-discrepancy sequence.
- ▶ QMC convergence is asymptotically much faster than MC convergence, but note that $\log(M)^d/M \gg M^{-1/2}$ for all reasonably sized **Figure:** F *M* even in fairly moderate dimensions *d*. E.g., numbers for *d* = 8 roughly from $M \ge 1.8 \times 10^{29}$.



Figure: Pairs of one-dimensional Sobol numbers





Randomized quasi Monte Carlo



Setting: $(x_i)_{i=1}^M$ low-discrepancy sequence in dim. d, $(U_l)_{l \in \mathbb{N}}$ i.i.d. sequence, $U_l \sim \mathcal{U}([0, 1]^d)$.

Definition (Randomized quasi Monte Carlo simulation)

$$J_{M;m}^{R}[f] \coloneqq \frac{1}{m} \sum_{l=1}^{m} \frac{1}{M} \sum_{i=1}^{M} f(\mathbf{x}_{i} + U_{l} \pmod{1})$$





Setting: $(x_i)_{i=1}^M$ low-discrepancy sequence in dim. d, $(U_l)_{l \in \mathbb{N}}$ i.i.d. sequence, $U_l \sim \mathcal{U}([0, 1]^d)$.

Definition (Randomized quasi Monte Carlo simulation)

$$J_{M;m}^{R}[f] \coloneqq \frac{1}{m} \sum_{l=1}^{m} \frac{1}{M} \sum_{i=1}^{M} f(x_{i} + U_{l} \pmod{1})$$

- Provides variance reduction for the Monte Carlo estimator $\frac{1}{m} \sum_{l=1}^{m} f(U_l)$
- Provides sharp, computable error control for the quasi Monte Carlo approximation $\frac{1}{M} \sum_{i=1}^{M} f(x_i):$ $E\left[\left(I[f] - J_{M;m}^R[f]\right)^2\right] = \frac{\operatorname{var}\left(\frac{1}{M} \sum_{i=1}^{M} f(x_i + U \pmod{1})\right)}{m}$

• Typically $m \ll M$, as accuracy is more important than error control.

