Stochastic collocation and MLMC methods for elliptic PDEs with random coefficients

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Acknowlegments: L. Tamellini, F. Tesei (EPFL), R. Tempone, (KAUST), J. Beck (UCL)

Workshop "Partial Differential Equations with Random Coefficients" WIAS, Berlin, 13 15 November, 2013



Italian project FIRB-IDEAS ('09) Advanced Numerical Techniques for Uncertainty Quantification in Engineering and Life Science Problems

Center for Advanced Modeling and Science





Outline

1 Model problem: elliptic PDEs with random coefficients

- 2 Polynomial approximation by sparse grid collocation
 - Quasi optimal sparse grid construction
 - Convergence result for elliptic PDEs with random inclusions
 - Numerical results

3 Combined MLMC / Sparse grid method

- MLMC with Sparse Grid Control Variate
- Variance analysis and algorithm tuning
- Numerical results

Conclusions

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Model problem: elliptic PDE with random coeffs

Let (Ω, \mathcal{F}, P) be a complete probability space and $D \subset \mathbb{R}^d$ and open bounded domain.

$$\begin{cases} -\operatorname{div}(a(\omega, x)\nabla u(\omega, x)) = f(x) & x \in D, \ \omega \in \Omega, \\ u(\omega, x) = 0 & x \in \partial D, \ \omega \in \Omega \end{cases}$$

with $f \in L^2(D)$ and $a(\omega, x) : \Omega \times D \to \mathbb{R}$ an almost surely bounded random field.

$$\|u(\omega,\cdot)\|_V \leq \frac{C_P}{a_{min}(\omega)} \|f\|_{L^2(D)},$$
 a.s. in Ω

Therefore, $u \in L^p_P(\Omega, V)$ for all $p \leq \bar{p}$. In particular, $u \in L^2_P(\Omega, \nabla)$



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Coercivity assumption: $a_{min}(\omega) = \operatorname{essinf}_{x \in D} a(\omega, x) > 0$ almost surely and $\mathbb{E}[a_{min}^{-\bar{p}}] < \infty$ for some $\bar{p} \geq 2$.

Then $u \in V = H_0^1(D)$ almost surely and $\|u(\omega, \cdot)\|_V \leq \frac{C_P}{1-|f||_{L^2(D)}} \|f\|_{L^2(D)}$, a.s. in

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Random field parametrization

In view of polynomial approximation we parametrize the random field $a(\omega, x)$ by a finite or countable sequence of random variables $\mathbf{y}(\omega) = (y_1(\omega), \ldots, y_N(\omega))$ with range $\Gamma = \mathbf{y}(\Omega) \subset \mathbb{R}^N$ and probability density function $\rho : \Gamma \to \mathbb{R}_+$:

 $a(\omega, x) = a(\mathbf{y}(\omega), x)$

Then the stochastic solution u depends on ω only through the vector $\mathbf{y}(\omega)$: $u(\omega, x) = u(\mathbf{y}(\omega), x)$

parameter-to-solution map: $u(\mathbf{y}) : \Gamma \to V$, $u \in L^2_o(\Gamma, V)$.



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Examples of random diffusion coefficients

Inclusions problem

Ν

y describes the conductivity in each inclusion



$$a(\mathbf{y}, \mathbf{x}) = a_0 + \sum_{n=N}^{\infty} y_n \mathbb{1}_{D_n}(\mathbf{x})$$

with $y_n \sim \mathcal{U}([y_{min}, y_{max}])$ and $y_{min} > -a_0$.

Therefore $a_{min}(\mathbf{y}) \in L^p_{\rho}(\Gamma)$ for any $1 \leq p \leq \infty$.

 $\implies u \in L^p_{\rho}(\Gamma, H^1_0(D)), \ \forall 1 \le p \le \infty$

Random fields problem

 $a(\mathbf{y}, \mathbf{x})$ is a random field, e.g. lognormal: $a(\mathbf{y}, \mathbf{x}) = e^{\gamma(\mathbf{y}, \mathbf{x})}$ with γ expanded e.g. in Karhunen-Loève series

 $\gamma(\mathbf{y}, \mathbf{x}) = \sum_{n=1}^{\infty} \sqrt{\lambda_n} y_n b_n(\mathbf{x}), \quad y_n \sim N$

If $Cov[\gamma]$ is Holder continuous, then $a_{min} \in L^{\rho}_{\rho}(\Gamma)$ for any $1 \leq p < \infty$ (see e.g. [Charrier, 2011])

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 $p_n(x), \quad y_n \sim N(0,1) \ i.i.d.$

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Stochastic multivariate polynomial approximation

 The parameter-to-solution map u(y) : Γ → V is often smooth (even analytic for the elliptic diffusion model). It is therefore sound to approximate it by global multivariate polynomials.

 Let Λ ⊂ ℕ^N be an index set of cardinality |Λ| = M, and consider the multivariate polynomial space

 $\mathbb{P}_{\Lambda}(\Gamma) = span\left\{\prod_{n=1}^{N} y_n^{p_n}, \quad \text{with } \mathbf{p} = (p_1, \dots, p_N) \in \Lambda\right\}$

We seek an approximation $P_{\Lambda} u \in \mathbb{P}_{\Lambda}(\Gamma) \otimes V$.

Collocation approaches

Construct a polynomial approximation of $u(\mathbf{y}) : \Gamma \to V$ using only point evaluations $u_i = u(\mathbf{y}_i)$ where $\{\mathbf{y}_i\}_{i=1}^{\tilde{M}}$ is a set of suitable collocation points, with $\tilde{M} \ge M$.

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- Let $\mathbf{i} = [i_1, \dots, i_N] \in \mathbb{N}_+^N$ and $m(i) : \mathbb{N}_+ \to \mathbb{N}_+$ an increasing function
 - 1D polynomial interpolant operators: $\mathscr{U}_n^{m(i_n)}$ on $m(i_n)$ abscissas. We use either
 - Clenshaw-Curtis (extrema on Chebyshev polynomials)
 - Gauss points w.r.t. the weight ρ_n , assuming that the probability density factorizes as $\rho(\mathbf{y}) = \prod_{n=1}^{N} \rho_n(y_n)$
 - 2 Detail operator: $\Delta_n^{m(i_n)} = \mathscr{U}_n^{m(i_n)} \mathscr{U}_n^{m(i_n-1)}, \ \mathscr{U}_n^{m(0)} = 0.$
 - **3** Hierarchical surplus: $\Delta^{m(i)} = \bigotimes_{n=1}^{N} \Delta_n^{m(i_n)}$.
 - Sparse grid approximation: on an index set $\mathcal{I} \subset \mathbb{N}^N$

$$\mathcal{S}_{\mathcal{I}} u = \sum_{\mathbf{i} \in \mathcal{I}} \Delta^{m(\mathbf{i})}[u]$$

Assumption: The set \mathcal{I} is downward closed: $\mathbf{i} \in \mathcal{I} \implies \mathbf{i} - \mathbf{e}_n \in \mathcal{I}, n = 1, \dots, N.$

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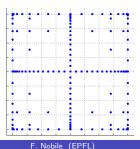


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Equivalent formulation

$$\begin{split} \mathcal{S}_{\mathcal{I}} u &= \sum_{\mathbf{i} \in \mathcal{I}} c(\mathbf{i}) \mathscr{U}_1^{m(i_1)} \otimes \dots \otimes \mathscr{U}_N^{m(i_N)} u. \\ \text{with} \quad c(\mathbf{i}) &= \sum_{\substack{\mathbf{j} \in \{0,1\}^N \\ (\mathbf{i}+\mathbf{j}) \in \mathcal{I}}} (-1)^{j_1 + \dots + j_N}, \quad \text{and} \ c(\mathbf{i}) = 0 \quad \text{if} \quad \mathbf{i} + \mathbf{1} \in \mathcal{I} \end{split}$$

• linear combination of tensor grids (each with relatively few points!)



Theorem ([Back-Nobile-Tamellini-Tempone, 2010]) Let $\Lambda(\mathcal{I}, m) = \{ \mathbf{p} \in \mathbb{N}^N : \mathbf{p} \le m(\mathbf{i}) - 1, \mathbf{i} \in \mathcal{I} \}.$ Then

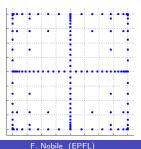
- $\mathcal{S}_{\mathcal{I}}: C^0(\Gamma) \to \mathbb{P}_{\Lambda(\mathcal{I},m)}(\Gamma)$
- $S_{\mathcal{I}}v = v, \quad \forall v \in \mathbb{P}_{\Lambda(\mathcal{I},m)}(\Gamma)$

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Equivalent formulation

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$$\mathcal{S}_{\mathcal{I}} u = \sum_{\mathbf{i} \in \mathcal{I}} \Delta^{m(\mathbf{i})}[u] \implies \|u - \mathcal{S}_{\mathcal{I}} u\| = \|\sum_{\mathbf{i} \notin \mathcal{I}} \Delta^{m(\mathbf{i})}[u]\| \le \sum_{\mathbf{i} \notin \mathcal{I}} \|\Delta^{m(\mathbf{i})}[u]\|$$

One can use a knapsack problem-approach [Griebel-Knapek '09, Gerstner-Griebel '03, Bungartz-Griebel '04] to select the best \mathcal{I} : for each multiindex i:

 Estimated error contribution (how much error decreases if i is added to I)

$\Delta E(\mathbf{i}) \geq \|\Delta^{m(\mathbf{i})}[u]\|_V$

 Estimated work contribution (how much the work, i.e. number of evaluations, increases if i is added to I)

 $\Delta W({f i})$ such that $\sum \Delta W({f i}) \geq W({\cal I})$

where $W(\mathcal{I})$ is the total number of points in the sparse \mathfrak{g}



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$$\Delta W(\mathbf{i}) \quad \text{such that} \quad \sum_{\mathbf{i} \in \mathcal{I}} \Delta W(\mathbf{i}) \geq W(\mathcal{I})$$

where $W(\mathcal{I})$ is the total number of points in the sparse grid

Then estimate the profit of each **i** as

$$P(\mathbf{i}) = rac{\Delta E(\mathbf{i})}{\Delta W(\mathbf{i})}$$

and build the sparse grid using the set \mathcal{I}_M of the M indices with the largest estimated profit.

 $\mathcal{I}_M := \{\mathbf{i} \in \mathbb{N}^N \mid P(\mathbf{i}) \geq P_M^{ord}\}$

where $\{P_i^{ord}\}_j$ is the ordered sequence of profits.

If the set \mathcal{I}_M is not downward closed (lower), take the smallest lower set $\tilde{\mathcal{I}}_M \supset \mathcal{I}_M$. This is equivalent to consider the modified profits $\tilde{P}(\mathbf{i}) = \max_{\mathbf{j} > \mathbf{i}} P(\mathbf{j})$. (see [Chkifa-Cohen-DeVore-Schwab M2AN '13])

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 A priori approach [Back-Nobile-Tamellini-Tempone '11]. Whenever possible, use *a-priori/a-posteriori* information to build the optimal set. Avoids the "exploration" cost that can be expensive in high dimension.

The estimate of $\Delta E(\mathbf{i})$ can be related to the decay of the coefficients of the gPC expansion $u = \sum_{\mathbf{p}} u_{\mathbf{p}}\psi_{\mathbf{p}}$ of the solution onto an orthonormal polynomial basis (Legendre, Chebyshev, Hermite, ...) and to the Lebesgue constant of the interpolation scheme.

 A posteriori approach [Gerstner-Griebel '03, Klimke, PhD '06]. Given a set Λ, explore all the neighbor multi-indices (margin) and pick up those corresponding to the largest profits.



• A priori approach [Back-Nobile-Tamellini-Tempone '11]. Whenever possible, use *a-priori/a-posteriori* information to build the optimal set. Avoids the "exploration" cost that can be expensive in high dimension.

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General convergence result

Theorem [Tamellini PhD thesis '12], [Nobile-Tamellini-Tempone, in preparation] Let $S_{\tilde{I}_{M}}u$ be the quasi-optimal sparse grid approximation and $W_{\tilde{I}_{M}}$ the total number of points in the sparse grid.

If
$$C(\tau) := \left(\sum_{\mathbf{i} \in \mathbb{N}^N} \tilde{P}(\mathbf{i})^{\tau} \Delta W(\mathbf{i})\right)^{\frac{1}{\tau}} < \infty$$
 for some $\tau < 1$

$$\Gamma \mathsf{hen} \qquad \|u - \mathcal{S}_{\widetilde{\mathcal{I}}_M} u\|_{L^2_\rho(\Gamma, V)} \leq C(\tau) \mathcal{W}_{\widetilde{\mathcal{I}}_M}^{1 - \frac{1}{\tau}}$$



The proof uses

• Stechkin lemma: given a non-negative descreasing sequence $\{a_k\}_k$, then

$$\sum_{k=N+1}^{\infty} a_k \leq N^{1-\frac{1}{\tau}} \left(\sum_{k=1}^{\infty} a_k^{\tau}\right)^{\frac{1}{\tau}}, \quad 0 < \tau < 1.$$

• ordered repeated sequence of profits $\{\hat{P}_k\}_k = \{\underbrace{\tilde{P}_1, \ldots, \tilde{P}_1}_{\Delta W_1 \text{ times}}, \underbrace{\tilde{P}_2, \ldots, \tilde{P}_2}_{\Delta W_2 \text{ times}}, \ldots\}.$

• Let $W_M = \sum_{j=1}^M \Delta W_j$ and observe that $W_M \ge W_{\tilde{\mathcal{I}}_M}$. Then

$$\begin{aligned} \|u - \mathcal{S}_{\tilde{\mathcal{I}}_{M}} u\|_{L^{2}_{p}(\Gamma, V)} &\leq \sum_{k=M+1}^{\infty} \Delta E_{k} \leq \sum_{k=W_{M}+1}^{\infty} \hat{P}_{k} \quad \text{[apply Stechkin]} \\ &\leq \|\{\hat{P}_{k}\}_{k}\|_{l^{r}} W_{M}^{1-\frac{1}{r}} \leq C(\tau) W_{\tilde{\mathcal{I}}_{M}}^{1-\frac{1}{r}} \end{aligned}$$

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How to estimate ΔW and ΔE

- △W(i): number of new points in ⊗^N_{n=1} Δ^{m(in)}_n
 Count all points in U^{m(i1)}₁ ⊗ · · · ⊗ U^{m(iN)}_N (non-nested case) or just the extra points added (nested case)
- ② $\Delta E(\mathbf{i})$: use expansion on a suitable basis $u = \sum_{\mathbf{p}} u_{\mathbf{p}} \psi_{\mathbf{p}}$ (Legendre, Chebyshev, ...) and relate $\Delta E(\mathbf{i})$ with the decay of the Fourier coefficients $u_{\mathbf{p}}$.

E.g. for Chebyshev expansion $(\|\psi_{\mathbf{p}}\|_{\infty} = 1)$

$$\Delta E(\mathbf{i}) \leq 2 \mathbb{L}_{m(\mathbf{i})} \sum_{\mathbf{p} \geq m(\mathbf{i}-1)} \|u_{\mathbf{p}}\|_{V},$$

where $\mathbb{L}_{m(i)} = \prod_{n=1}^{N} \mathbb{L}_{m(i_n)}$ and $\mathbb{L}_{m(i_n)} := \|\mathscr{U}_n^{m(i_n)}\|_{\mathcal{L}(C^0, L^2_{\rho})}$ is the Lebesgue constant form C^0 to L^2_{ρ} .

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How to estimate ΔW and ΔE

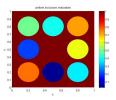
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Elliptic equation with random inclusions



• $Y_i \sim \mathcal{U}[a_i, b_i]$, independent

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$$\Gamma = \prod_{i=1}^{N} [a_i, b_i]$$

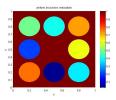
•
$$u(\mathbf{y}): \Gamma \to H^1_0(D)$$

- The solution $u(\mathbf{y}): \Gamma \to H^1_0(D)$ is analytic in a polydisk in the complex plane \mathbb{C}^N
- The solution can be expanded in Chebyshev series
 u(**y**) = ∑_{**p**} *u*_{**p**}ψ_{**p**}(**y**) with ||ψ_{**p**}||_∞ ≤ 1. Estimates on Chebyshev
 coefficients are available [Babuska-Nobile-Tempone '07]
 [Nobile-Tamellini-Tempone '13]

$$\|u_{\mathbf{p}}\|_{H^{1}_{0}(D)} \leq Ce^{-\sum_{n=1}^{N}g_{i}p_{i}}$$



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Quasi-optimal sparse grid construction using (nested) Clenshaw-Curtis points

- $m(i) = 2^{i-1} + 1$ (doubling the points). $\Delta W(\mathbf{i}) \approx \prod_{n=1}^{N} 2^{i_n-2}$
- Decay of Chebyshev coefficients: $\|u_p\|_V \leq Ce^{-g\sum_{n=1}^N p_n}$
- Lebesgue constant: $\mathbb{L}(i) \leq \frac{2}{\pi} \log(i+1) + 1$
- Error estimate: $\Delta E(\mathbf{i}) = C \prod_{n=1}^{N} e^{-\hat{g}m(i_n-1)}$ Profit estimate: $P(\mathbf{i}) = \tilde{C}e^{-\hat{g}\sum_{n=1}^{N}m(i_n-1)}$



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Theorem [Nobile-Tamellini-Tempone '13]

$$\begin{split} \left(\sum_{\mathbf{i} \in \mathbb{N}^N} P(\mathbf{i})^{\tau} \Delta W(\mathbf{i}) \right)^{\frac{1}{\tau}} < \infty \quad \text{for all } 0 < \tau < 1 \\ \implies \quad \| u - \mathcal{S}_{\tilde{\mathcal{I}}_M} u \|_{L^2_\rho(\Gamma, V)} \leq C(\tau) \mathcal{W}_{\tilde{\mathcal{I}}_M}^{1 - \frac{1}{\tau}}, \quad \forall \, 0 < \tau < 1 \end{split}$$

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By optimizing with respect to τ one can get the estimate

$$\|u - \mathcal{S}_{\tilde{\mathcal{I}}_{M}} u\|_{L^{2}_{\rho}(\Gamma, V)} \leq C_{1}(N) \exp\{-NC_{2} W_{\tilde{\mathcal{I}}_{M}}^{\tilde{N}}\}$$



Quasi-optimal sparse grid construction using (non-nested) Legendre points

- m(i) = i (adding 1 point at the time). $\Delta W(\mathbf{i}) = \prod_{n=1}^{N} i_n$
- Decay of Chebyshev coefficients: $||u_p||_V \leq Ce^{-g\sum_{n=1}^N p_n}$
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$$\|u - \mathcal{S}_{\tilde{\mathcal{I}}_{M}} u\|_{L^{2}_{\rho}(\Gamma, V)} \leq \tilde{C}_{1}(N) \exp\{-N\tilde{C}_{2} W^{\frac{1}{2N}}_{\tilde{\mathcal{I}}_{M}}\}$$

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We have the theoretical bound $||u_{\mathbf{p}}||_{V} \approx Ce^{-\sum_{n=1}^{N} g_{n}p_{n}}$, hence, in particular, for $\mathbf{p} = \mathbf{e}_{j} = (0, \dots, 0, 1, 0, \dots, 0)$, $||u_{\mathbf{e}_{j}}||_{V} \approx Ce^{-g_{j}p_{j}}$ which implies that a 1D polynomial interpolation in the variable y_{j} only converges exponentially with rate g_{j} .

- The rates g_j can be estimated numerically by "1D analyses". (increase the polynomial degree in one variable at the time and fit the convergence rate).
- Once the rates g_i are available, we build the quasi optimal index set based on the profit estimate

$$P(\mathbf{i}) = rac{\Delta E(\mathbf{i})}{\Delta W(\mathbf{i})} \propto \prod_{n=1}^{N} rac{\mathbb{L}_{m(i_n)} e^{-g_{i_n} m(i_n-1)}}{\Delta W(i_n)}$$



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When working with random fields instread of inclusion problmes a good estimate for the Legendre/Chebyshev coefficients is (see [Bech-N.-Tamellini-Tempone M3AS '12, Cohen-DeVore-Schwab FoCM '10])

$$\|u_{\mathbf{p}}\|_{V} \approx C \frac{|\mathbf{p}|!}{\mathbf{p}!} \exp\{-\sum_{n=1}^{N} g_{n} p_{n}\}.$$

The rates g_n can be estimated again by 1D inexpensive analyses.



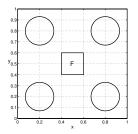
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$$\|u_{\mathbf{p}}\|_{V} \approx C \frac{|\mathbf{p}|!}{\mathbf{p}!} \exp\{-\sum_{n=1}^{N} g_{n} \rho_{n}\}.$$

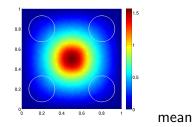
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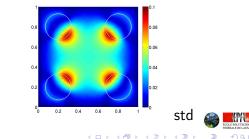
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Isotropic test case - 4 random inclusions



- Conductivity coefficient: matrix k=1 circular inclusions: k|_{Ω_i} ~ U(0.01, 1.99) → 4 iid uniform random variables
- forcing term $f = 100 \mathbb{1}_F$
- zero boundary conditions
- quantity of interest $\psi(u) = \int_F u$





Isotropic test case - 4 random inclusions

convergence plot for $\|\psi(u) - \mathcal{S}_{\mathcal{I}}\psi(u)\|_{L^{2}_{o}(\Gamma)}$ versus # pts sparse grid

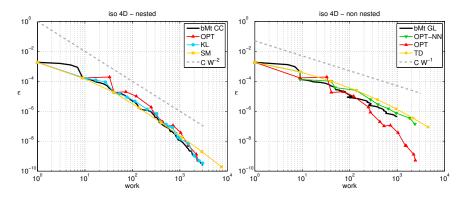


Figure: Results for the isotropic problem. Left: (nested) CC points Right: (non-nested) Gauss-Legendre points



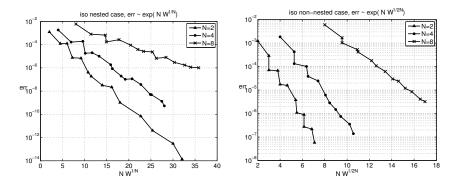
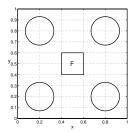


Figure: Results for the isotropic problem. Optimal sparse grids and their predicted convergence rates. **Top**: Nested CC points. **Bottom**: Non-nested Gauss Legendre points.



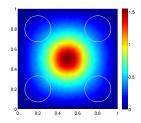
Anisotropic test case – 4 random inclusions

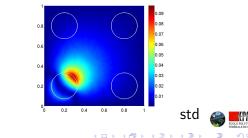
mean



• Conductivity coefficient: matrix k=1 circular inclusions: $k|_{\Omega_i} \sim \gamma_i \mathcal{U}(-0.99, 0.99) \rightarrow 2$ iid uniform random variables

•
$$\gamma_{1,2,3,4} = 1, 0.06, 0.0035, 0.0002$$





Anisotropic test case - 4 random inclusions

convergence plot for $\|\psi(u) - \mathcal{S}_{\mathcal{I}}\psi(u)\|_{L^{2}_{\rho}(\Gamma)}$ versus # pts sparse grid

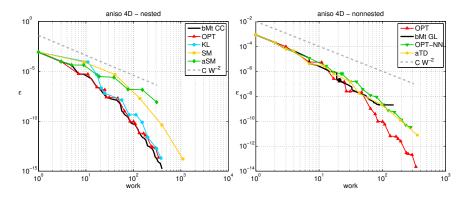


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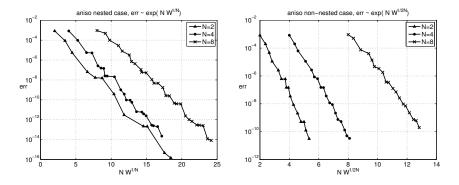
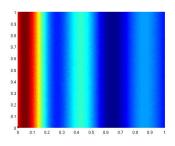


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Numerical test - 1D stationary lognormal field



 $L = 1, D = [0, L]^2.$

$$\begin{cases} -\nabla \cdot a(\mathbf{y}, \mathbf{x}) \nabla u(\mathbf{y}, \mathbf{x}) = 0\\ u = 1 \text{ on } x = 0, \ h = 0 \text{ on } x = 1\\ \text{no flux otherwise} \end{cases}$$

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$$egin{aligned} & \mathbf{a}(\mathbf{x},\mathbf{y})=e^{\gamma(\mathbf{x},\mathbf{y})},\ \mu_{\gamma}(\mathbf{x})=\mathbf{0},\ & \mathcal{C}ov_{\gamma}(\mathbf{x},\mathbf{x}')=\sigma^{2}e^{-rac{|\mathbf{x}_{1}-\mathbf{x}_{1}'|^{2}}{LC^{2}}} \end{aligned}$$

We approximate γ as

$$\gamma(\mathbf{y}, \mathbf{x}) \approx \mu(\mathbf{x}) + \sigma a_0 y_0 + \sigma \sum_{k=1}^{K} a_k \left[y_{2k-1} \cos\left(\frac{\pi}{L} k \mathbf{x}_1\right) + y_{2k} \sin\left(\frac{\pi}{L} k \mathbf{x}_1\right) \right]$$

with $y_i \sim \mathcal{N}(0, 1)$, i.i.d.
Given the Fourier series $\sigma^2 e^{-\frac{|z|^2}{LC^2}} = \sum_{k=0}^{\infty} c_k \cos\left(\frac{\pi}{L} k z\right)$, $a_k = \sqrt{c_k}$.





- Quantity of interest: effective permeability $\mathbb{E}[\Phi(u)]$, with $\Phi = \left[\int_0^L k(\cdot, x) \frac{\partial u(\cdot, x)}{\partial x} dx\right]$
- Convergence: $|\mathbb{E}[\Phi(\mathcal{S}_{\tilde{\mathcal{I}}_{M}}u)] \mathbb{E}[\Phi(u)]|$
- We compare Monte Carlo estimate with quasi-optimal sparse grids based on Gauss-Hermite-Patterson points (nested Gauss-Hermite)
- Estimate of Hermite coefficients decay:
 - for the simpler problem $\nabla \cdot a(\mathbf{y}) \nabla u(\mathbf{y}, \mathbf{x}) = f$, $a(\mathbf{y}) = e^{b_0 + \sum_{n=1}^{N} y_n b_n}$, we have $\|u_{\mathbf{i}}\|_V = C \frac{b_n^{i_n}}{\sqrt{i_n!}}$.
 - Heuristic: use the same ansatz $||u_i||_V \approx C \prod_{n=1}^N \frac{e^{-g_n i_n}}{\sqrt{i_n!}}$ but estimate the rates g_n numerically.



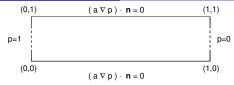


- Quantity of interest: effective permeability $\mathbb{E}[\Phi(u)]$, with $\Phi = \left[\int_0^L k(\cdot, x) \frac{\partial u(\cdot, x)}{\partial x} dx\right]$
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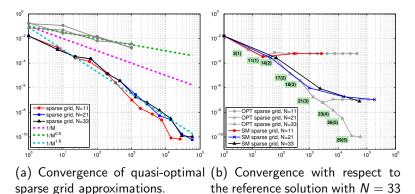
Polynomial approximation by sparse grid collocation Numerical results



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 - Heuristic: use the same ansatz $||u_i||_V \approx C \prod_{n=1}^N \frac{e^{-g_n i_n}}{\sqrt{i_n!}}$ but estimate the rates g_n numerically.

F. Nobile (EPFL)

Correlation length: LC = 0.2, Std: $\sigma = 0.3$ (c.o.v. $\sim 30\%$)



r.vs.

- The quasi optimal construction automatically adds new variables when needed.
- No need to truncate a-priori the random field

"A quasi-optimal sparse grids procedure for groundwater flows" by J. Beck, F. Nobile, a Tamellini and R. Tempone. To appear, LNCSE, Springer, 2013.



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 - inclusion type problems with few inclusions
 - smooth random fields with long correlation length
- On the other hand, these techniques suffer in the cases of
 - rough fields even with long correlation lenght; e.g. exponential covariance: $Cov_a(x, y) = \sigma^2 e^{|x-y|/l_c}$

$$a(\omega, x) = \overline{a}(x) + \sum_{n=1}^{\infty} \sqrt{\lambda_n} y_n(\omega) b_n(x), \qquad \lambda_n \sim n^{-2}$$

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 $\lambda_{n}=O(1)$ for $n\leq diam(D)/l_{c}$

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Outline

1 Model problem: elliptic PDEs with random coefficients

- 2 Polynomial approximation by sparse grid collocation
 - Quasi optimal sparse grid construction
 - Convergence result for elliptic PDEs with random inclusions
 - Numerical results

3 Combined MLMC / Sparse grid method

- MLMC with Sparse Grid Control Variate
- Variance analysis and algorithm tuning
- Numerical results

Conclusions

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Darcy Problem with log-normal permeability

find $\mathbf{u}(x,\omega): \overline{D} \times \Omega \to \mathbb{R}^d$ and $p(x,\omega): \overline{D} \times \Omega \to \mathbb{R}$ such that almost everywhere in $\omega \in \Omega$ it holds:

$$\begin{cases} \mathbf{u}(x,\omega) = -\mathbf{a}(x,\omega)\nabla p(x,\omega) & \text{in } D, \\ \operatorname{div}(\mathbf{u}(x,\omega)) = f(x) & \text{in } D, \\ + \text{ boundary conditions} & \text{on } \partial D. \end{cases}$$

- D is the bounded physical domain and (Ω, F, P) is the probability space
- **u** and *p* are the Darcy velocity and pressure, respectively
- a(x, ω) is the permeability field modeled as a lognormally distributed random field a(x, ω) = e^{γ(x,ω)}
- $\gamma(x,\omega)$ is a Gaussian stationary random field



Matérn Covariance Function

Matérn Family

$$\operatorname{cov}_{\gamma}(x_{1} - x_{2}) = \frac{\sigma^{2}}{\Gamma(\nu)2^{\nu-1}} \left(\sqrt{2\nu} \frac{\|x_{1} - x_{2}\|}{L_{c}}\right)^{\nu} K_{\nu}\left(\sqrt{2\nu} \frac{\|x_{1} - x_{2}\|}{L_{c}}\right)$$

- L_c is a correlation length
- Γ is the gamma function
- K_{ν} is the modified Bessel function of the second kind

•
$$\nu = 0.5$$
: $\operatorname{cov}_{\gamma}(x_1 - x_2) = \sigma^2 e^{-\frac{\|x_1 - x_2\|}{L_c}}$ (exponential covariance)
• $\nu \to \infty$: $\operatorname{cov}_{\gamma}(x_1 - x_2) = \sigma^2 e^{-\frac{\|x_1 - x_2\|^2}{L_c^2}}$ (Gaussian covariance)

Matérn Covariance Function: Regularity issues

Let $\nu = s + \alpha$ with $s = \lceil \nu - 1 \rceil \in \mathbb{N}$ and $\alpha = \nu - \lceil \nu - 1 \rceil \in (0, 1]$. Then the realizations of the random field are almost surely Hölder continuous, $\gamma(x, \omega) \in \mathcal{C}^{s,\beta}(\overline{D})$ with $\beta < \alpha$.

- For ν = 0.5 the covariance function is only Lipschitz continuous and the field is almost surely Hölder continuous γ(x,ω) ∈ C^{0,α}(D̄) with α < 0.5.
- For ν → ∞ the covariance function as well as the field are continuous with all their derivates, namely cov_γ(x), γ(x,ω_i) ∈ C[∞](D̄) ∀ω ∈ Ω

SEE Graham, Kuo, Nichols, Scheichl, Schwab, Sloan '13



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MLMC

Q = Q(p): Qol related to the solution of the PDE $\mathcal{T}_{h_0}, ..., \mathcal{T}_{h_L}$: sequence of increasingly fine triangulations $Y_{h_l} = Q_{h_l} - Q_{h_{l-1}}$: difference of the Qol between two consecutive grids.

Telescopic sum + Linearity of expectation:

$$\mathbb{E}[Q_{h_L}] = \sum_{l=0}^{L} \mathbb{E}[Y_{h_l}], \quad Q_{h_{-1}} = 0$$

MLMC Estimator:

$$\hat{Q}_{L,\{M_l\}}^{MLMC} = \sum_{l=0}^{L} rac{1}{M_l} \sum_{i=1}^{M_l} \left(Q_{h_l}(\omega_i) - Q_{h_{l-1}}(\omega_i)
ight)$$
ntrup, Scheichl, Giles, Ullmann '12], [Charrier, Scheichl, Teckentrup '11]

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MLMC – Mean Square Error (MSE)

$$e(\hat{Q}_{L,\{M_l\}}^{MLMC})^2 = \underbrace{\sum_{l=0}^{L} \frac{\mathbb{V}ar(Y_{h_l})}{M_l}}_{(i)} + \underbrace{\mathbb{E}[Q_{h_L} - Q]^2}_{(ii)}$$

 M_l : number of samples on each level. A good choice of M_l , for l = 0, ..., L, represents a crucial issue for the effectiveness of the method (see [Cliffe, Giles, Scheichl, Teckentrup, '11], [Barth, Schwab, Zollinger, '11]) (*i*): represents the variance of the estimator, i.e. the statistical error:

it is expected to be significantly smaller than the variance of the standard MC estimator

(*ii*): represents the bias of the error, due to the finite element (FE) discretization

Control Variate

Idea: use the solution of an auxiliary problem with regularized coefficient as control variate Problem: we do not know exactly the expected value of the control variate. However this can be computed efficiently by a Stochastic Collocation (SC) technique

Original Problem Auxiliary Problem

- $\begin{cases} -\operatorname{div}(a\nabla p) = f & \text{in } D, \\ + \text{ boundary conditions } & \text{on } \partial D. \end{cases} \begin{cases} -\operatorname{div}(a^{\epsilon}\nabla p^{\epsilon}) = f & \text{in } D, \\ + \text{ boundary conditions } & \text{on } \partial D. \end{cases}$
 - $a = e^{\gamma}$: random field obtained starting from the Matérn covariance function with parameter ν

•
$$a^{\epsilon} = e^{\gamma^{\epsilon}}$$
: regularized version of a



Control Variate

Regularized Gaussian random field obtained via convolution with a Gaussian kernel

$$\gamma^{\epsilon} = \gamma * \phi_{\epsilon}(x), \text{ where } \phi_{\epsilon} = rac{1}{(2\pi\epsilon^2)^{rac{d}{2}}} e^{-rac{\|x\|^2}{2\epsilon^2}}$$

Quantity of interest defined via control variate

 $Q^{CV} = Q - (Q^{\epsilon} - \mathbb{E}[Q^{\epsilon}])$

where Q = Q(p) and $Q^{\epsilon} = Q(p^{\epsilon})$

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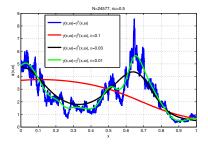
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0.5

Control Variate



Lipschitz continuous covariance function (u = 0.5)

Twice differentiable covariance function ($\nu = 2.5$)

N=24577, nu=2.5



 $\gamma(x,\omega)=\gamma^{V}(x,\omega)$ $\gamma(x,\omega)=\gamma^{F}(x,\omega), \epsilon=0.1$

(x.ω)=γ^ε(x.ω), ε=0.01

MLMC with Control Variate

 $\mathcal{T}_{h_0}, ..., \mathcal{T}_{h_L}$: sequence of increasingly fine triangulations; $Y_{h_l}^{CV} = Q_{h_l}^{CV} - Q_{h_{l-1}}^{CV}$: difference of the Qol between two consecutive grids.

Telescopic sum + Linearity of expectation:

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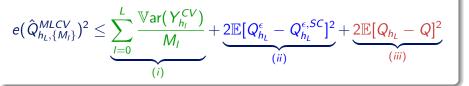
MLCV Estimator

$$\hat{Q}_{h_L,\{M_l\}}^{MLCV} = \sum_{l=0}^{L} rac{1}{M_l} \sum_{i=1}^{M_l} \left(Q_{h_l}^i - Q_{h_{l-1}}^i - (Q_{h_l}^{\epsilon,i} - Q_{h_{l-1}}^{\epsilon,i})
ight) + \mathbb{E}[Q_{h_L}^{\epsilon,SC}]$$

Variance analysis and algorithm tuning

MLMC with Control Variate

Mean Square Error (MSE)



 $\mathbb{E}[Q_{h_{\ell}}^{\epsilon,SC}]$: mean of the Qol $Q_{h_{\ell}}^{\epsilon}$ computed with a SC scheme on sparse grids.

(*i*): variance of the estimator

(ii): bias due to the SC approximation of the mean of the control variate $\mathbb{E}[Q^{\epsilon}]$

(*iii*): bias due to the finite element approximation

Remark: if ϵ tends to 0 the statistical error (i) vanishes; on the statistical error (i) va hand keeping small the SC error (ii) becomes too costly.

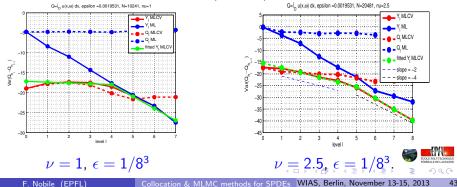
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Error estimate (preliminary result)

$$\mathbb{E}\left[\left(Q-Q_{h_l}-\left(Q^{\epsilon}-Q_{h_l}^{\epsilon}\right)\right)^2\right]^{\frac{1}{2}} \leq ch_l^{\min\{\nu,p\}} \min_{s=1,...,p} h_l^{\min\{\nu,s\}} \epsilon^{\min\{(\nu-s)_+,2\}}$$

Variance of the difference of the Qol between consecutive grids. The dashed lines represent the slopes h_1^2 and h_1^4 .



MLCV Algorithm

- **1** Given a prescribed *tol* select h_L in such a way to have (*iii*) $\leq tol^2$;
- Set h₀ = O(|D|) and evaluate (iv) = Var(Y^{CV}_{h_l}) and (v) = Var(Q^{CV}_{h_l}); we can select among two basic strategies: Strategy 1: if (iv) < (v) ∀ I apply the MLCV scheme starting from level 0;
 - Strategy 2: if $(iv) \approx (v)$ for $l = 0, ..., l_0$ set $h_0 = h_{l_0}$; and use the control variate only on level l_0 and a standard MLMC on subsequent levels, namely

$$\hat{Q}_{h_{L},\{M_{l}\}}^{MLCV} = \frac{1}{M_{l_{0}}} \sum_{i=1}^{M_{l_{0}}} \left(Q_{h_{l_{0}}}^{i} - Q_{h_{l_{0}}}^{\epsilon,i} \right) + \sum_{l=l_{0}+1}^{L} \frac{1}{M_{l}} \sum_{i=1}^{M_{l}} \left(Q_{h_{l}}^{i} - Q_{h_{l-1}}^{i} \right) + \mathbb{E}[Q_{h_{l_{0}}}^{\epsilon,SC}]$$

3 according to the strategy selected compute the number of samples *M_l* for *l* = 0, ..., *L* and the number of knots of the sparse grid *M_{SG}* by solving an optimization problem in such a way to have

 (i) + (ii) ≤ tol²
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 F. Nobile (EPFL)
 Collocation & MLMC methods for SPDEs WIAS, Berlin, November 13-15, 2013

Optimization Problem

- ϵ is considered fixed
- C_l is the computational cost needed to solve one deterministic system on the grid of mesh size h_l ; cost model: $C_l = 2C_{l-1} = ... = 2^l C_0 = \gamma 2^l h_0^{-1}$. For instance see [Cliffe, Giles, Scheichl, Teckentrup, '11]

Computational cost

- strategy 1: $C(M_l, M_{SG}) = 2M_0C_0 + 2\sum_{l=1}^{L} M_l(C_l + C_{l-1}) + M_{SG}C_L$
- strategy 2: $C(M_l, M_{SG}) = 2M_{l_0}C_{l_0} + \sum_{l=l_0+1}^{L} M_l(C_l + C_{l-1}) + M_{SG}C_{l_0}$ Associated Error fitted model
 - strategy 1: $e(M_I, M_{SG}) = \sum_{l=0}^{L} \frac{\min\left\{c_1 h_l^{2\min\{\nu, p\}}, c_2 h_l^{4\min\{\nu, p\}}\right\}}{M_l} + c_3 M_{SG}^{\alpha}$

• strategy 2: $e(M_I, M_{SG}) = \sum_{I=I_0}^{L} \frac{c_1 h_I^{4 \min\{\nu, p\}}}{M_I} + c_2 M_{SG}^{\alpha}$

Perform a Lagrange optimization by considering

$$\mathcal{L}(M_I, M_{SG}, \lambda) = C(M_I, M_{SG}) - \lambda(e(M_I, M_{SG}) - tol^2)$$

Optimization Problem

Values obtained (strategy one):

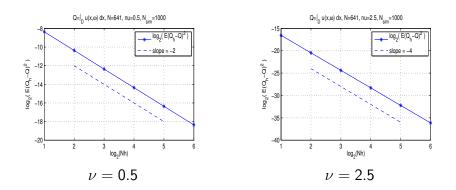
•
$$M_0 = \sqrt{-\lambda} \sqrt{\frac{c_1 h_0^{2 \min\{\nu, \rho\}}}{2C_0 2^l}} = \sqrt{-\lambda} \sqrt{\frac{V_0}{2C_0}}$$

• $M_l = \sqrt{-\lambda} \sqrt{\frac{\min\{c_1 h_l^{2 \min\{\nu, \rho\}}, c_2 h_l^{4 \min\{\nu, \rho\}}\}}{3C_0 2^l}} = \sqrt{-\lambda} \sqrt{\frac{V_l}{3C_0 2^l}}$ for $l = 1, ..., L$
• $M_{SG} = (-\lambda)^{\frac{1}{1-\alpha}} \left(-\alpha 2^{-L} \frac{c_3}{C_0}\right)^{\frac{1}{1-\alpha}}$

where λ has to be computed from:

$$\frac{1}{\sqrt{-\lambda}} = \frac{tol^2}{\sqrt{C_0}(\sqrt{2\nu_0} + \sum_{l=1}^L \sqrt{3\nu_l}2^l) + c_3(\frac{1}{\sqrt{-\lambda}})^{\frac{-1-\alpha}{1-\alpha}}(-\alpha 2^{-L}\frac{c_3}{C_0})^{\frac{\alpha}{1-\alpha}}}$$

Approximation error $\mathbb{E}[Q_{h_l} - Q]^2$



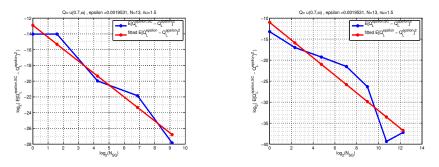


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$$\mathsf{SC} \; \mathsf{error} \; \mathbb{E}[Q_{h_l}^{\epsilon,\mathsf{SC}}-Q^\epsilon]^2$$



 $\nu = 0.5$; fitted slope $\alpha = -1.5$ $\nu = 2.5$; fitted slope $\alpha = -2.2$

In both cases more than the 99% of the variability has been taken into account. The fitted rate is better than MC in both cases.

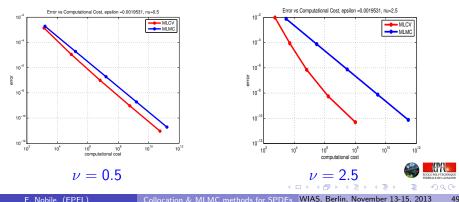


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Error vs Cost : MLCV vs MLMC

Error and computational cost associated to several values of $tol = 10^{-1}, ..., 10^{-5}$. $\epsilon = 1/8^3$ in both cases.



error = statistical error + SC error

F. Nobile (EPFL)

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4 Conclusions

- We have analyzed the convergence of quasi-optimal sparse grid approximations based on the selection of the M most profitable hierarchical surpluses. Convergence rates are related to summability properties of the profits, weighted by the corresponding works.
- Sharp a-priori / a-posteriori analysis of the decay of the polynomial chaos expansion of the solution allows to construct optimized sparse grids that provide effective approximations also in infinite dimensions for smooth fields.
- The "profit based" a-posteriori adaptive algorithm is also performing very efficiently, close to the best approximation.
- for rough, longly correlated fields, a good idea is to use a sparse grid polynomial approximation on a smoothed problem as a control variate in a MLMC algorith. Preliminary results show a considerable improvement of the overall complexity.

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Thank you for your attention!



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